

Statistical Decision Theory with Counterfactual Loss

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 - What happens if we use *all potential outcomes*?
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Contribution: Extend classical decision theory for treatment choice to counterfactual losses.

Statistical Decision Theory

Wald 1950: Decision-making as a game against nature.

- ① Nature picks an unknown state θ ,
- ② Decision-maker chooses action $D = d$,
- ③ A **loss** $\ell(d, \theta)$ quantifies the cost of choosing d under θ .

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Given covariates \mathbf{X} , construct a decision rule $D = \pi(\mathbf{X})$.

Measure performance with **risk**,

$$R(\pi; \theta, \ell) = \mathbb{E}_\theta [\ell(\pi(\mathbf{X}), \theta)].$$

Treatment Choice

Manski [2000; 2004; 2011]: Statistical decision theory for treatment choice.

Idea:

- Choose treatment $D = d$ to minimize loss based on outcome Y .
- Loss depends on potential outcome $Y(d)$, i.e., $\ell(d, Y(d))$.

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Idea:

- Choose treatment $D = d$ to minimize loss based on outcome Y .
- Loss depends on potential outcome $Y(d)$, i.e., $\ell(d, Y(d))$.

Given covariates \mathbf{X} , use a treatment rule $D = \pi(\mathbf{X})$.

Evaluate risk

$$R(\pi; \ell) = \mathbb{E} [\ell(\pi(\mathbf{X}), Y(\pi(\mathbf{X})))]$$

Limitation: Loss only depends on the treated potential outcome.

Trichotomous Decision

Physician treating a patient

- $D = 0$: No treatment
- $D = 1$: Standard treatment (more invasive)
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c_D cost of treatment D

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- $Y = 1$: survival
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- $(D, Y(D)) = (1, 0) : \ell_0 + c_1$
- $(D, Y(D)) = (2, 1) : \ell_1 + c_2$

Trichotomous Decision II

Standard Loss: $\ell^{\text{STD}}(D, Y(D)) = \ell_{Y(D)} + c_D$

Clinical & ethical goal: Avoid overtreatment

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- Prefer option $k < d$ if $Y(k) = 1$
- r_k regret of overtreating option k

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$$(Y(0), Y(1), Y(2)) = (0, 1, 1)$$

- $D = 1 : \ell_1 + c_1$
- $D = 2 : \ell_1 + c_2 + r_1$

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We show:

- For r_k sufficiently large, ℓ^{STD} and ℓ^{COF} yield different treatment preferences.
- No standard loss that can take these ethical considerations into account.

Setup

Observed data: For each unit $i = 1, \dots, n$, observe (\mathbf{X}_i, D_i, Y_i) , where:

- Covariates: $\mathbf{X}_i \in \mathcal{X}$
- Decision: $D_i \in \mathcal{D} = \{0, 1, \dots, K - 1\}$
- Outcome: $Y_i \in \mathcal{Y} = \{0, 1, \dots, M - 1\}$
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Assumptions:

- **IID Sampling:** $\{Y_i, D_i, D_i^*, \mathbf{X}_i\}$ are IID
- **Consistency:** $Y_i = Y_i(D_i)$, and if $D_i^* = D_i$, then $Y_i(D_i^*) = Y_i(D_i)$
- **Strong Ignorability:**
 - *Unconfoundedness:* $D_i \perp\!\!\!\perp (D_i^*, \{Y_i(d)\}_{d \in \mathcal{D}}) \mid \mathbf{X}_i$
 - *Overlap:* $\exists \eta > 0 : \eta < \Pr(D_i = d \mid \mathbf{X}_i) < 1 - \eta$, for all $d \in \mathcal{D}$

Counterfactual Loss and Risk

Counterfactual loss: $\ell : \mathcal{D} \times \mathcal{Y}^K \times \mathcal{X} \rightarrow \mathbb{R}$, i.e., $\ell(d; y_1, \dots, y_K, \mathbf{x})$.

Loss of choosing $D^* = d$ given

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Given counterfactual loss ℓ , the counterfactual risk of decision D^* is:

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Problem: $\Pr(D^* = d, Y(0) = y_0, \dots, Y(K-1) = y_{K-1} \mid \mathbf{X} = \mathbf{x})$ **unidentifiable**



Identifiability of Counterfactual Risk

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Can we impose structure on ℓ that enables identification?

Additivity

Definition (Additive Counterfactual Loss)

Let $\mathbf{y} = (y_0, \dots, y_{K-1}) \in \mathcal{Y}^K$. Then the additive counterfactual loss is defined as,

$$\ell^{\text{ADD}}(d; \mathbf{y}, \mathbf{x}) = \omega_d(d, y_d, \mathbf{x}) + \sum_{k \in \mathcal{D}, k \neq d} \omega_k(d, y_k, \mathbf{x}) + \varpi(\mathbf{y}, \mathbf{x}).$$

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- $\omega_d(d, y_d, \mathbf{x}) : \mathcal{D} \times \mathcal{Y} \times \mathcal{X} \rightarrow \mathbb{R}$
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 - $\omega_k(d, y_k) = r_k y_k \mathbb{1}\{k < d\}$
- $\varpi(\mathbf{y}, \mathbf{x}) : \mathcal{Y}^K \times \mathcal{X} \rightarrow \mathbb{R}$
 - Intercept term
 - Decision independent
 - $\varpi(\mathbf{y}) = 0$

Additivity Implies Identifiability

Theorem (Additivity Implies Identifiability)

Let ℓ^{ADD} be additive. Then,

$$R(D^*; \ell^{\text{ADD}}) = \sum_{d \in \mathcal{D}} \sum_{k \in \mathcal{D}} \sum_{y \in \mathcal{Y}} \mathbb{E}[\omega_k(d, y, \mathbf{x}) \Pr(D^* = d, Y(k) = y \mid \mathbf{X})] + \mathbb{E}[C(\mathbf{X})],$$

where

$$C(\mathbf{x}) = \sum_{\mathbf{y} \in \mathcal{Y}^K} \varpi(\mathbf{y}, \mathbf{x}) \Pr(\mathbf{Y}(\mathcal{D}) = \mathbf{y} \mid \mathbf{X} = \mathbf{x}),$$

with $\mathbf{Y}(\mathcal{D}) = (Y(0), \dots, Y(K-1))$.

Decomposition into identifiable marginal term and unidentifiable term, not depending on D^* .

Thus an additive loss yields an identifiable risk (up to a constant).

Additivity is Necessary and Sufficient

Can a counterfactual risk be identified under a non-additive loss?

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Theorem

Under strong ignorability, the counterfactual risk $R(D^; \ell)$ is identifiable (up to a constant) if and only if the loss ℓ is additive.*

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- Binary outcome
- Connections between loss and principal strata
- When additive counterfactual losses yield different treatment recommendations than standard losses

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Next steps and extensions:

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- Relaxing strong ignorability
- Continuous outcomes Y

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Thank you!

Happy to talk counterfactuals: What should I have done?
Scan the QR code to view the paper.



Binary Outcomes

Corollary

Assume $Y \in \{0, 1\}$. Let ℓ^{ADD} be additive. Then,

$$\begin{aligned} R_x(D^*; \ell^{\text{ADD}}) &= \sum_{d \in \mathcal{D}} \zeta_d(d, x) \Pr(D^* = d, Y(d) = 1 \mid \mathbf{X} = x) \\ &+ \sum_{d \in \mathcal{D}} \sum_{k \in \mathcal{D}, k \neq d} \zeta_k(d, x) \Pr(D^* = d, Y(k) = 1 \mid \mathbf{X} = x) \\ &+ \sum_{d \in \mathcal{D}} \xi(d, x) \Pr(D^* = d \mid \mathbf{X} = x) + C(x). \end{aligned}$$

where $\zeta_k(d, x) = \omega_k(d, 1, x) - \omega_k(d, 0, x)$, $\xi(d, x) = \sum_{k \in \mathcal{D}} \omega_k(d, 0, x)$, and

$$C(x) = \sum_{\mathbf{y} \in \{0,1\}^K} \varpi(\mathbf{y}, x) \Pr(\mathbf{Y}(\mathcal{D}) = \mathbf{y} \mid \mathbf{X} = x).$$

Decomposition into **accuracy**, **difficulty**, **decision** and **unidentifiable** constant term.

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Choosing weights

$$\omega_d(d, 1) \leq \{\omega_{d'}(d, 0)\}_{d' \neq d} \leq 0 \leq \{\omega_{d'}(d, 1)\}_{d' \neq d} \leq \omega_d(d, 0),$$

yields $\zeta_d(d, x) \leq 0 \leq \zeta_k(d, x)$ for all $k \neq d$. Implies risk decreases with accuracy and increases with counterfactual regret.

Proposition (Additive Counterfactual Risk with Binary Decision)

Suppose that the decision is binary, i.e., $\mathcal{D} = \{0, 1\}$. For any additive counterfactual loss $\ell^{\text{ADD}}(d; \mathbf{y}, \mathbf{x})$, we can construct a standard loss $\ell^{\text{STD}}(d, y_d)$ such that the risk difference $R(D^*; \ell^{\text{ADD}}) - R(D^*; \ell^{\text{STD}})$ does not depend on D^* .

If decisions are binary, any additive counterfactual risk admits a standard risk that yields identical treatment recommendations.

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However,

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- ℓ^{STD} has infinitely many additive counterfactual losses ℓ^{ADD} 's with the same treatment recommendations
- Each of them assigns different values to principal strata
- Thus ℓ^{STD} has no clear interpretation based on principal strata
- While ℓ^{ADD} does have a clear interpretation based on principal strata

Proposition (Additive Counterfactual Risk with Non-binary Decision)

Assume that the decision is non-binary, i.e., $K = |\mathcal{D}| \geq 3$. Then, for any additive counterfactual loss with at least one counterfactual weight $\omega_k(d, y_k, \mathbf{x})$ depending on the decision $d \in \mathcal{D}$ and potential outcome $y_k \in \mathcal{Y}$ for $d \neq k$, there exists no standard loss $\ell^{\text{STD}}(d; y_d)$ such that the risk difference $R(D^; \ell^{\text{ADD}}) - R(D^*; \ell^{\text{STD}})$ does not depend on D^* .*

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Counterfactual loss (extended from Ben-Michael, Greiner, et al. 2024):

$$\ell(D; Y(0), Y(1)) = \ell_{Y(D)} + \tilde{\ell}_{Y(1-D)} + c_D$$

$\tilde{\ell}_y$ loss of counterfactual outcome, $\tilde{\ell}_0 < \tilde{\ell}_1$, i.e. loss is greater when the patient survived under the other treatment (missed positive)

Example: Asymmetric Counterfactual Loss

Hippocratic Oath — “Do no harm”: Causing harm with treatment is worse than failing to provide treatment.

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Loss based on Principal Strata (Ben-Michael, Imai, and Jiang 2024):

$$\begin{aligned}\ell(D; Y(0), Y(1)) = & (1 - Y(0))Y(1)\ell_D^R + Y(0)(1 - Y(1))\ell_{1-D}^H \\ & + Y(0)Y(1)\ell_1 + (1 - Y(0))(1 - Y(1))\ell_0 + c_D,\end{aligned}$$

Example: Asymmetric Counterfactual Loss

Hippocratic Oath — “Do no harm”: Causing harm with treatment is worse than failing to provide treatment.

Loss based on Principal Strata (Ben-Michael, Imai, and Jiang 2024):

$$\ell(D; Y(0), Y(1)) = (1 - Y(0))Y(1)\ell_D^R + Y(0)(1 - Y(1))\ell_{1-D}^H + Y(0)Y(1)\ell_1 + (1 - Y(0))(1 - Y(1))\ell_0 + c_D,$$

Asymmetry in loss:

$$\underbrace{\Delta^R = \ell_0^R - \ell_1^R}_{\text{Failure to treat a responder}} < \underbrace{\Delta^H = \ell_0^H - \ell_1^H}_{\text{Harming a patient}}.$$

Non-additive loss if $\Delta^R \neq \Delta^H$.

References I

-  Ben-Michael, Eli, D. James Greiner, et al. (Mar. 2024). *Does AI help humans make better decisions? A methodological framework for experimental evaluation*. arXiv:2403.12108 [cs, econ, q-fin, stat]. DOI: 10.48550/arXiv.2403.12108. URL: <http://arxiv.org/abs/2403.12108> (visited on 07/03/2024).
-  Ben-Michael, Eli, Kosuke Imai, and Zhichao Jiang (2024). "Policy Learning with Asymmetric Counterfactual Utilities". In: *Journal of the American Statistical Association* 0.0. Publisher: ASA Website _eprint: <https://doi.org/10.1080/01621459.2023.2300507>, pp. 1–14. ISSN: 0162-1459. DOI: 10.1080/01621459.2023.2300507. URL: <https://doi.org/10.1080/01621459.2023.2300507> (visited on 09/20/2024).
-  Manski, Charles F (2000). "Identification problems and decisions under ambiguity: Empirical analysis of treatment response and normative analysis of treatment choice". In: *Journal of Econometrics* 95.2, pp. 415–442.
-  — (2004). "Statistical treatment rules for heterogeneous populations". In: *Econometrica* 72.4, pp. 1221–1246.

References II

-  Manski, Charles F (2011). "Choosing treatment policies under ambiguity". In: *Annu. Rev. Econ.* 3.1, pp. 25–49.
-  Wald, Abraham (1950). *Statistical decision functions*. Wiley.